

HOMEWORK SET 2  
*Martingales, stopping times (mostly)*  
 Martingale Theory with Applications, 1<sup>st</sup> teaching block, 2017  
 School of Mathematics, University of Bristol

Problems with •'s are to be handed in. These are due in class or in the blue locker with my name on the ground floor of the Main Maths Building before 16:00pm on Friday, 20<sup>th</sup> October. Please show your work leading to the result, not only the result. Each problem worth the number of •'s you see right next to it. Hence, for example, Problem 2.3 worth four marks.

2.1 *A few missing bits from the proof of Scheffé's Theorem.*

- (a) Show that for any non-negative reals  $x_n$  and  $x$ ,  $(x_n - x)^- \leq x$ .
- (b) Show that for any reals  $x_n$  and  $x$ ,  $|x_n^+ - x^+| \leq |x_n - x|$ .
- (c) Show that for any reals  $x_n$  and  $x$ ,  $|x_n^- - x^-| \leq |x_n - x|$ .
- (d) Show, however, that for any reals  $x_n$  and  $x$ ,  $|x_n - x| \leq |x_n^+ - x^+| + |x_n^- - x^-|$ .

2.2 Let  $\mathcal{G} \subset \mathcal{F}$  be a  $\sigma$ -algebra on the probability space  $(\Omega, \mathcal{F}, \mathbb{P})$ . If  $X$  is  $\mathcal{G}$ -measurable, then  $\mathbb{E}(X | \mathcal{G}) = X$  which suggests that the map  $X \mapsto \mathbb{E}(X | \mathcal{G})$  is a projection. Show that indeed: this map is an orthogonal projection in the Hilbert space  $\mathcal{L}^2(\Omega, \mathcal{F}, \mathbb{P})$  (with inner product  $\langle X, Y \rangle_{\mathbb{P}} = \mathbb{E}(XY)$ ) onto the subspace  $\mathcal{L}^2(\Omega, \mathcal{G}, \mathbb{P})$ .

2.3 ••• Let  $\xi_1, \xi_2, \dots$  be iid. standard normal random variables. (Recall their moment generating function:  $\mathbb{E}(e^{\lambda \xi_i}) = e^{\lambda^2/2}$ .) Let  $a, b \in \mathbb{R}$ ,

$$S_n = \sum_{k=1}^n \xi_k, \quad \text{and} \quad X_n = e^{aS_n - bn}.$$

Show that

$$X_n \rightarrow 0 \text{ a.s.} \Leftrightarrow b > 0,$$

but for any  $r \geq 1$

$$X_n \rightarrow 0 \text{ in } \mathcal{L}^r \Leftrightarrow r < \frac{2b}{a^2}.$$

2.4 Let  $S$  and  $T$  be stopping times w.r.t. the filtration  $\mathcal{F}_n$ . Which of these are stopping times? Explain.

$$S \wedge T := \min(S, T), \quad S \vee T := \max(S, T), \quad T + S, \quad T - S \text{ (assume } T \geq S \text{ here.)}$$

2.5 ••• Let  $S_n$  be a simple symmetric random walk on the square lattice  $\mathbb{Z}^2$  with  $S_0 = (0, 0)$ . That is, the walker starts from the origin and at each step independently, she steps one unit to East, North, West or South with equal chance. Denote by  $D_n$  the walker's distance from the origin of  $\mathbb{Z}^2$  at time  $n$ , and let  $\nu_r = \inf\{n : D_n > r\}$ .

- (a) Show that  $D_n^2 - n$  is a martingale.
- (b) Show that  $r^{-2} \mathbb{E} \nu_r \rightarrow 1$  as  $r \rightarrow \infty$ .

2.6 The problem is the same as the previous one, except that the walk is on  $\mathbb{R}^2$  and steps are of length one in iid. Uniform(0,  $2\pi$ ) directions.

2.7 Let  $X_1, X_2, \dots$  be iid. Exponential(1) random variables,  $S_n = X_1 + \dots + X_n$ , and  $\{\mathcal{F}_n\}$  the natural filtration. Show that

$$\frac{n!}{(1 + S_n)^{n+1}} e^{S_n}$$

is a martingale w.r.t.  $\{\mathcal{F}_n\}$ .

2.8 •••• An urn contains  $n$  white and  $n$  black balls. We draw them one by one without replacement. We pay £1 for any black ball drawn but receive £1 for any white one. Denote by  $X_i$  our money after the  $i^{\text{th}}$  draw ( $X_0 = 0$ ). Let

$$Y_i = \frac{X_i}{2n - i} \quad (1 \leq i \leq 2n - 1), \quad \text{and} \quad Z_i = \frac{X_i^2 - (2n - i)}{(2n - i)(2n - i - 1)} \quad (1 \leq i \leq 2n - 2).$$

(a) Show that both  $Y_i$  and  $Z_i$  are martingales.

(b) Calculate the variance of  $X_i$ .

2.9 Let  $X_j, j \geq 1$ , be absolutely integrable random variables, and  $\mathcal{F}_n := \sigma(X_j, 1 \leq j \leq n)$ ,  $n \geq 0$ , their natural filtration. Define the new random variables

$$Z_0 := 0, \quad Z_n := \sum_{j=0}^{n-1} (X_{j+1} - \mathbb{E}(X_{j+1} | \mathcal{F}_j)).$$

Prove that the process  $n \mapsto Z_n$  is an  $(\mathcal{F}_n)_{n \geq 0}$ -martingale.

2.10 A biased coin shows HEAD with probability  $\theta \in (0, 1)$ , and TAIL with probability  $1 - \theta$ . The value  $\theta$  of the bias is *not known*. For  $t \in [0, 1]$  and  $n \in \mathbb{N}$  we define  $p_{n,t} : \{0, 1\}^n \rightarrow [0, 1]$  by

$$p_{n,t}(x_1, x_2, \dots, x_n) = t^{\sum_{j=1}^n x_j} \cdot (1 - t)^{n - \sum_{j=1}^n x_j}.$$

We make two hypotheses about the possible value of  $\theta$ : either  $\theta = a$ , or  $\theta = b$ , where  $a, b \in [0, 1]$  and  $a \neq b$ . We toss the coin repeatedly and form the sequence of random variables

$$Z_n := \frac{p_{n,a}(\xi_1, \xi_2, \dots, \xi_n)}{p_{n,b}(\xi_1, \xi_2, \dots, \xi_n)},$$

where we write  $\xi_j = 1$  if the  $j^{\text{th}}$  flip is HEAD and  $\xi_j = 0$  if it is TAIL. Show that the process  $n \mapsto Z_n$  is a martingale (w.r.t. the natural filtration generated by the coin tosses) if and only if the true bias of the coin is  $\theta = b$ .

2.11 Let  $\eta_n$  be a homogeneous Markov chain on the countable state space  $S := \{0, 1, 2, \dots\}$  and  $\mathcal{F}_n := \sigma(\eta_j, 0 \leq j \leq n)$ ,  $n \geq 0$  its natural filtration. For  $i \in S$  denote by  $Q(i)$  the probability that the Markov chain starting from site  $i$  ever reaches the point  $0 \in S$ :

$$Q(i) := \mathbb{P}\{\exists m < \infty : \eta_m = 0 \mid \eta_0 = i\}.$$

Prove that  $Z_n := Q(\eta_n)$  is an  $(\mathcal{F}_n)_{n \geq 0}$ -martingale.

2.12 *Galton-Watson Branching Process*. Let  $\xi_{n,k}$ ,  $n = 1, 2, \dots$ ,  $k = 1, 2, \dots$  be iid.  $\mathbb{N}$ -valued random variables with finite mean  $\mu$  and variance  $\sigma^2$ . Define the Galton-Watson branching process

$$Z_0 := 1, \quad Z_{n+1} := \sum_{k=1}^{Z_n} \xi_{n+1,k}$$

and let  $\mathcal{F}_n := \sigma(Z_j : 0 \leq j \leq n)$ ,  $n \geq 0$ , be its natural filtration.

- (a) Prove that  $M_n := \mu^{-n} Z_n$ ,  $n \geq 0$ , is an  $(\mathcal{F}_n)_{n \geq 0}$ -martingale.
- (b) Prove that  $\mathbb{E}(Z_{n+1}^2 | \mathcal{F}_n) = \mu^2 Z_n^2 + \sigma^2 Z_n$ .
- (c) Using the result from (b) prove that

$$N_n := \begin{cases} M_n^2 - \frac{\sigma^2}{\mu^{n+1}} \frac{\mu^n - 1}{\mu - 1} M_n, & \text{if } \mu \neq 1, \\ M_n^2 - n\sigma^2 M_n, & \text{if } \mu = 1 \end{cases}$$

is also an  $(\mathcal{F}_n)_{n \geq 0}$ -martingale.

- (d) Using the result from (c) prove that if  $\mu > 1$  then  $\sup_{0 \leq n < \infty} \mathbb{E} M_n^2 < \infty$  (that is: the martingale  $M_n$  is uniformly bounded in  $\mathcal{L}^2$ ) while if  $\mu \leq 1$  then  $\lim_{n \rightarrow \infty} \mathbb{E} M_n^2 = \infty$ .

2.13 We repeatedly toss a fair coin.

- (a) What is the expected number of tosses until we have seen the pattern HTHT for the first time?
- (b) What is the expected number of tosses until we have seen the pattern THTH for the first time?
- (c) What is the expected number of tosses until we have seen the pattern HTTH for the first time?
- (d) What is the expected number of tosses until we have seen the pattern THHT for the first time?
- (e) Give an example of a four letter pattern of H-s and T-s that has the maximal expected number of tosses, of any four letter patterns, until it is seen.

2.14 ••• The previous question with a biased coin. Explain your answer.

2.15 Let  $m \geq 2$  be an integer. At time  $n = 0$ , an urn contains  $2m$  balls of which  $m$  are red and  $m$  are blue. At each time  $n = 1, 2, \dots, 2m$  we draw a randomly chosen ball without replacement from the urn and record its colour. For  $n = 0, 1, \dots, 2m - 1$  let  $N_n$  denote the number of red balls left in the urn after time  $n$ , and

$$P_n := \frac{N_n}{2m - n}$$

denote the fraction of them. Let  $(\mathcal{F}_n)_{0 \leq n \leq 2m}$  be the natural filtration generated by the process  $(N_n)_{0 \leq n \leq 2m}$ .

- (a) Show that  $n \mapsto P_n$  is an  $\mathcal{F}_n$ -martingale.
- (b) Let  $T$  be the first time at which the ball drawn is red. Show that the  $(T + 1)^{\text{st}}$  draw is equally likely to be red or blue.