

# Stability of multi-dimensional Markov chains, with applications

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# Network

- $M$  devices
- Wi-Max (IEEE 802.16) and Wi-Fi (IEEE 802.11) protocols
- Schedule for Wi-Max traffic
- Random access for Wi-Fi traffic
- Data streams of the same standard interfere with each other
- Data streams from different standards generated by different devices do not interfere
- Data streams from different standards generated by the same device do interfere

# Model

- $M$  transmitters
- Slotted time
- Each transmission duration equals 1
- High- and low- priority messages
- At most one high- and at most one low-priority messages transmitted in a time slot
- A transmitter cannot transmit both high- and low-priority messages in a time slot
- Schedule for high-priority messages, no collisions
- Random access for low-priority messages, collided messages return to their origins

# High-priority (red) messages assumptions

- At time slot  $t$ ,  $\xi_n^t$  new ones arrive at device  $n$ ,  $\mathbf{E}\xi_n^t = \lambda_R/M$
- $\{\xi_n^t\}$  are i.i.d. in  $t$
- Symmetrical schedule for transmissions: at time slot  $t$ , transmitter number  $i(t) = ((t - 1) \bmod M) + 1$  is scheduled to transmit a red message
- If transmitter  $i(t)$  has a red message, its transmission is attempted and is successful
- Otherwise no red message is transmitted in the time slot

# Red-messages dynamics

Let  $R_n^t$  be the number of red messages in the queue of transmitter  $n$  at time  $t$ . Then

$$(R_1^t, \dots, R_M^t)$$

is a Markov chain which is positive recurrent if  $\lambda_R < 1$ . Moreover,

$$\mathbf{P}(R_{i(t)}^t = 0) \rightarrow 1 - \lambda_R, \quad t \rightarrow \infty.$$

## Low-priority (green) messages assumptions

- At time slot  $t$ ,  $\eta_n^t$  new ones arrive at device  $n$ ,  $\mathbf{E}\eta_n^t = \lambda_G/M$
- $\{\eta_n^t\}$  are i.i.d. in  $t$
- Transmission attempts are governed by a random-access ALOHA-type protocol: every transmitter not currently transmitting a red message with a non-empty green queue attempts to transmit a green message with (fixed) probability  $p$
- Three possibilities:
  - No transmission attempted
  - Exactly one transmission attempted. It is successful
  - More than one transmission attempts. All of them unsuccessful, messages stay in their queues

# Model dynamics

Let  $G_n^t$  be the number of green messages in the queue of transmitter  $n$  at time  $t$ . Then

$$(G_1^t, \dots, G_M^t, R_1^t, \dots, R_M^t)$$

is a Markov chain and we are interested in its long-time behaviour.

Let first  $M = 1$ . Then a green message transmission will be attempted (and will always be successful) with probability  $p$  every time the red queue is empty. We expect that

$$\lambda_G < (1 - \lambda_R)p$$

leads to stability (positive recurrence).

# Model dynamics

Let now  $M \geq 2$ . Two cases:

- Transmitter  $i(t)$  has a red message to transmit. Then probability of a successful transmission of a green message is  $(M - 1)p(1 - p)^{M-2}$
- Transmitter  $i(t)$  does not have a red message to transmit. Then probability of a successful transmission of a green message is  $Mp(1 - p)^{M-1}$

Stability should be achieved if

$$\lambda_G < \lambda_R(M - 1)p(1 - p)^{M-2} + (1 - \lambda_R)Mp(1 - p)^{M-1}.$$



# General mathematical model

Let  $\{X^t\}$  and  $\{Y^t\}$  be random sequences taking values in measurable spaces  $(\mathcal{X}, \mathcal{B}_\mathcal{X})$  and  $(\mathcal{Y}, \mathcal{B}_\mathcal{Y})$ , respectively, and assume that  $\{(X^t, Y^t)\}$  is a Markov Chain. Assume also that

- $\{X^t\}$  is a Markov Chain with autonomous dynamics: for all  $x \in \mathcal{X}$  and  $y \in \mathcal{Y}$ ,

$$\mathbf{P}_{x,y}(X^1 \in \cdot) = \mathbf{P}_x(X^1 \in \cdot).$$

- The  $X$ -chain is *Harris ergodic*, so there exists a stationary distribution  $\pi = \pi_X$  such that

$$\sup_{B \in \mathcal{B}} |\mathbf{P}_x(X^t \in B) - \pi(B)| \rightarrow 0$$

as  $t \rightarrow \infty$ , for any  $x \in \mathcal{X}$ .

## Auxiliary chain

Introduce an auxiliary (time-homogeneous) Markov chain  $\{\hat{Y}^t\}$  with transition probabilities

$$\mathbf{P}(\hat{Y}^{t+1} \in \cdot \mid \hat{Y}^t = y) = \int_{\mathcal{X}} \pi_X(dx) \mathbf{P}(Y^1 \in \cdot \mid X^1 = x, Y^0 = y).$$

### Hypothesis

If  $\{\hat{Y}^t\}$  is positive recurrent, then so is  $(X^t, Y^t)$ .

### Theorem

*Natural (Foster-Lyapunov type conditions) imply positive recurrence of both  $\{\hat{Y}^t\}$  and  $(X^t, Y^t)$ .*

## Conditions on $Y$

For the sequence  $\{Y^t\}$  we assume that there exists a non-negative measurable function  $L_2$  such that:

- The expectations of the absolute values of the increments of the sequence  $\{L_2(Y^t)\}$  are bounded from above by a constant  $U$ :

$$\sup_{x \in \mathcal{X}, y \in \mathcal{Y}} \mathbf{E}_{x,y} |L_2(Y^1) - L_2(Y^0)| \leq U < \infty.$$

- There exist a non-negative and non-increasing function  $h(N)$ ,  $N \geq 0$  such that  $h(N) \downarrow 0$  as  $N \rightarrow \infty$ , and a measurable function  $f : \mathcal{X} \rightarrow (-\infty, \infty)$  such that

$$\int_{\mathcal{X}} f(x) \pi(dx) := -\varepsilon < 0$$

and

$$\mathbf{E}_{x,y} (L_2(Y^1) - L_2(y)) \leq f(x) + h(L_2(y))$$

for all  $x \in \mathcal{X}$  and  $y \in \mathcal{Y}$ .

# Main result

## Remark

Conditions on  $Y$  imply positive recurrence of  $\hat{Y}^t$ .

## Theorem (Foss, S, Turlikov (2012))

*Under assumptions for  $X$  and  $Y$ , the Markov chain  $\{(X^t, Y^t)\}$  is positive recurrent.*

## Corollary

Stability of the communication network under natural conditions

## (Some of the) related models

- *Bin-packing problems* (Gamarnik, 2004; Gamarnik and Squillante, 2005)
- *Cat-and-mouse Markov chain* (Litvak and Robert, 2012)
- *Other communication-networks applications* (Borst, Jonckheere, Leskela, 2008; Shah, Shin, 2012)
- Queueing systems, storage processes, etc.....

# Work in progress and research plans

- No knowledge of stationary distribution
- Non-autonomous dynamics, stronger dependence
- Other standard methods for proving stability
- Other applications
- Unstable first component

# Stochastic recursive sequences

Under some (very weak) assumptions, every Markov chain may be represented as a *stochastic recursive sequence*

$$X_{t+1} = f(X_t, \xi_t)$$

with a certain function  $f$  and an i.i.d. sequence  $\{\xi_t\}$ .

SRS (with specific functions  $f$ ) have applications in economics.

# Precautionary savings model (Huggett, 1993)

Agents maximise expected discounted utility

$$\mathbf{E} \left[ \sum_{t=0}^{\infty} \beta^t u(c_t) \right]$$

subject to a budget constraint

$$c_t + R^{-1} a_{t+1} \leq a_t + e_t,$$

a non-negativity constraint on consumption  $c_t \geq 0$  where  $e_t$  is the endowment,  $a_t$  is current assets,  $a_{t+1}$  are assets next period,  $c_t$  is consumption,  $\beta \in (0, 1)$  is the discount factor and  $R^{-1}$  is the price of assets.



# Precautionary savings model

The individual's maximisation problem can be written recursively as

$$v_e(a) = \max_{(c, a') \in \Gamma(a, e)} u(c) + \beta E_{e'|e} v_{e'}(a')$$

where

$$\Gamma(a, e) = \{(c, a') \mid c + R^{-1}a' \leq a + e, a' \geq \underline{a}, c \geq 0\}$$

is the constraint set and  $v_e(a)$  are the value functions (one for each realisation  $e$ ). The policy functions are  $c = c(a, e)$  and  $a' = f(a, e)$ , and  $f$  is continuous and non-decreasing in the first argument.

# Known results

We have

$$a_{t+1} = f(a_t, e_t), \quad \text{or} \quad X_{t+1} = f(X_t, \xi_t)$$

with a continuous and non-decreasing  $f$ .

*Question: does  $\{X_t\}$  converge to a stationary regime?*

- Most known results in economics assume  $\{\xi_t\}$  to be i.i.d.
- Huggett(1993) assumes that  $\{e_t\}$  is a two-state Markov chain with a positive correlation between  $e_t$  and  $e_{t+1}$ .

# Known results

## Theorem (Bhattacharya, Majumdar (2007))

Assume a time-homogeneous Markov chain  $\{X_t\}$  is represented as a stochastic recursion with i.i.d. driving sequence  $\{\xi_n\}$ , where function  $f : [0, 1] \times \mathcal{V} \rightarrow [0, 1]$  is monotone increasing in the first argument. Assume there exist a number  $c \in [0, 1]$  and an integer  $N \geq 1$  such that

$$\varepsilon_1 := \mathbf{P}^{(1)}(X_N \leq c) > 0$$

and

$$\varepsilon_2 := \mathbf{P}^{(0)}(X_N \geq c) > 0.$$

Then there exists a unique distribution  $\pi$  on  $[0, 1]$  such that

$$\sup_x d(F_t^{(x)}, \pi) \rightarrow 0, \quad n \rightarrow \infty$$

exponentially fast.

# Our model

Introduce an SRS

$$X_{t+1} = f(X_t, Z_t)$$

with a regenerative sequence  $\{Z_t\}$  and a non-decreasing and continuous (in the first argument) function  $f$ .

Introduce also an auxiliary process  $\tilde{X}_n^{(a)}$  that starts from  $\tilde{X}_0^{(a)} = a$  at time 0 and follows recursion

$$\tilde{X}_{n+1}^{(a)} = f\left(\tilde{X}_n^{(a)}, Z_{T_0+n}\right) \quad \text{for all } n \geq 0.$$

# Main result

## Theorem (Foss, S., Thomas, Worrall (2014))

Assume that the function  $f$  is monotone increasing in the first argument and the following assumptions hold:

$$\varepsilon_1 := \mathbf{P} \left( \tilde{X}_{T_1 - T_0}^{(1)} \leq c \right) > 0,$$

and

$$\varepsilon_2 := \mathbf{P} \left( \tilde{X}_{T_1 - T_0}^{(0)} \geq c \right) > 0.$$

Then there exists a distribution  $\tilde{\pi}$  on  $[0, 1]$  such that  $\sup_x d(G_n^{(x)}, \tilde{\pi}) \rightarrow 0$  as  $n \rightarrow \infty$  exponentially fast. Here  $G_n^{(x)}$  is the distribution of  $X_{T_n}$  if  $X_{T_0} = x$ .

Further, if in addition the function  $f$  is continuous in the first argument, then there exists a distribution  $\pi$  such that the distributions of  $X_t$  converge weakly to  $\pi$ , for any initial value  $X_0$ .

# Corollaries

- Stationary regime in the Huggett model with an arbitrary finite state space Markov chain
- Stationary regime in a number of other well-known economics models with Markov driving sequences rather than i.i.d.